



Facilitating probit likelihood optimization

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12 September 2023







- 1 The multinomial probit model: purpose and estimation
- 2 Numerical optimization and the initialization effect
- 3 Our initialization strategy for probit likelihood optimization
- 4 How does the strategy perform in comparison to random initialization?
- 5 Takeaways







1 The multinomial probit model: purpose and estimation

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One of the most important questions (for me) when attending a conference ...





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This is a discrete choice setting:

- deciders choose among a discrete set of alternatives
- based on decider- and/or alternative-specific attributes.





- politics (Ampel vs. opposition vs. non-voter)
- marketing (Apple vs. Samsung vs. ...)
- transportation (private vehicle vs. public transport vs. bike vs. ...)





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The multinomial probit model is widely used to analyze such discrete choices in fields like

- politics (Ampel vs. opposition vs. non-voter)
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- What is the probability of each choice being made?
- Which attributes greatly influence decisions?
- How are attribute trade-offs made?
- Do different population segments choose differently?
- What is the impact of new choice alternatives?





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Formally, the model

- connects attributes *X_n*
- to the choice $y_n \in \{1, \ldots, J\}$
- via latent utilities U_n that are defined as
 - 1. a linear function $V_n = X_n \beta$
 - 2. plus a Gaussian error ε_n
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In formulas:

 $U_n = V_n + \varepsilon_n$ $V_n = X_n\beta$ $\varepsilon_n \sim N(0, \Sigma)$ $y_n = \arg \max U_n$





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Under this model, the choice probability for alternative $j \in \{1, \ldots, J\}$ equals

$$P_{n,j} = \Phi(-\Delta_j X_n \beta \mid 0; \Delta_j \Sigma \Delta'_j).$$

 \Im Technical detail: the operator Δ_j takes utility differences with respect to alternative *j* for identification. Lennart Oelschläger Dietmar Bauer | Facilitating probit likelihood optimization



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This involves evaluating the Gaussian CDF, which has no closed form.

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In formulas:

 $U_n = V_n + \varepsilon_n$ $V_n = X_n\beta$ $\varepsilon_n \sim N(0, \Sigma)$ $y_n = \arg \max U_n$







How to fit this model to observed choice data $(X_n, y_n)_n$?



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1. define the log-likelihood function

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$$\ell(\beta, \Sigma \mid X, y) = \sum_{n,j} \mathbb{1}(y_n = j) \log P_{n,j}$$

2. and maximize it numerically over the parameters (β, Σ) .

♀ Technical detail: Σ is a restricted matrix but can be parametrized as a vector. Lennart Oelschläger Dietmar Bauer | Facilitating probit likelihood optimization



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- \blacksquare curse of dimensionality \Rightarrow optimization becomes slow
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Research question: How to make probit estimation faster and more reliable?

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1 The multinomial probit model: purpose and estimation

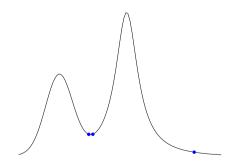
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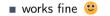








Depending on θ_0 , the optimization







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- 🔳 works fine 🙂
- 🔳 does not work 🙁



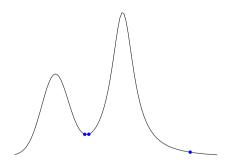


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Aim to initialize close to the global optimum

- 1. to reduce computation time and
- 2. to avoid local optima.







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Initialization strategy

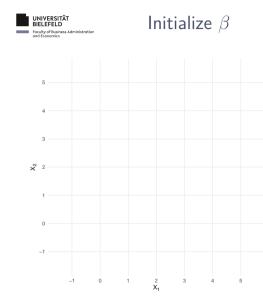


For the numerical optimization of the probit log-likelihood function

$$\ell(\beta, \Sigma \mid X, y) = \sum_{n,j} \mathbb{1}(y_n = j) \log P_{n,j}$$

we can quickly find consistent initial values

- β_0 via exploiting a constant utility direction,
- and Σ_0 conditional on β_0 via Gibbs sampling.

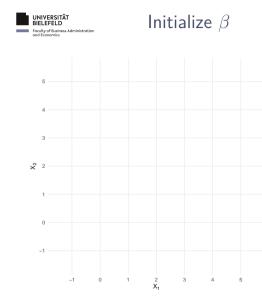




Assume

$$\begin{pmatrix} V_1 \\ V_2 \end{pmatrix} = \begin{pmatrix} X_1 & 0 \\ 0 & X_2 \end{pmatrix} \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}$$

(such a definition is typical for alternative-varying X with alternative-specific $\beta,$ e.g., travel time)





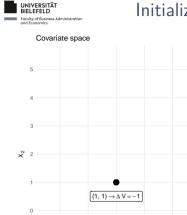
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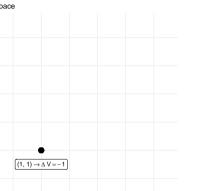


_1

-1

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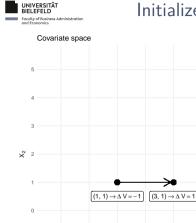
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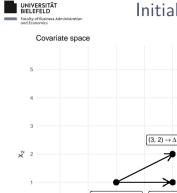
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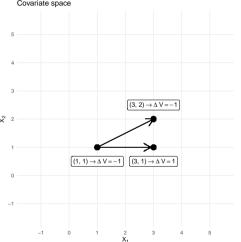
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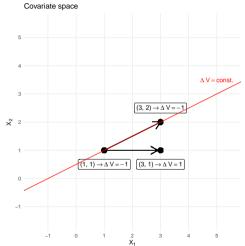
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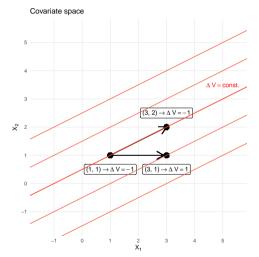
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We find the direction

 $\overrightarrow{\begin{pmatrix}1\\0.5\end{pmatrix}} = \overrightarrow{\begin{pmatrix}1/\beta_1\\1/\beta_2\end{pmatrix}}$







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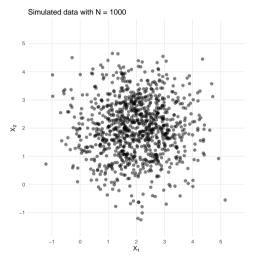
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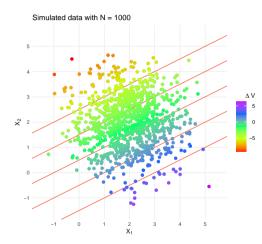
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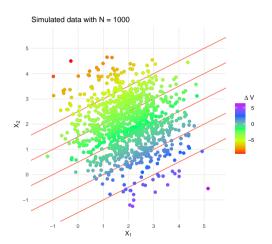








(depends on the unknown β)

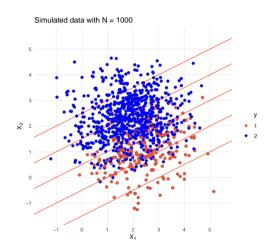


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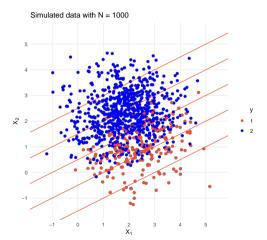


(depends on the unknown eta)

but we observe the choices y.



Initialize β



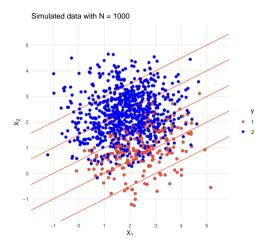


(depends on the unknown β) but we observe the choices γ .

They are disturbed by the error-term $\varepsilon.$ (here I used $\Sigma=\begin{pmatrix}2&1\\1&2\end{pmatrix}$)



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But we can still identify

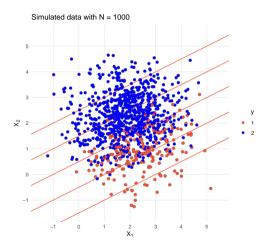
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as the kernel of Cor(X, y).

(constant choice probability in this direction)



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This gives an initial estimator $\hat{\beta}_0$ that can be shown to <u>consistent</u> as $N \to \infty$.



Initialize Σ



Now that we have a guess β_0 , we can draw Σ conditional on β_0 :

- 1. $(U_n)_n \mid \Sigma, \beta_0, (X_n, y_n)_n \sim \text{truncated normal}$
- 2. $\Sigma \mid \beta_0, (U_n)_n \sim \text{inverse Wishart}$



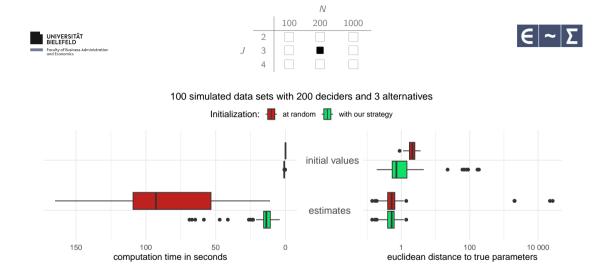


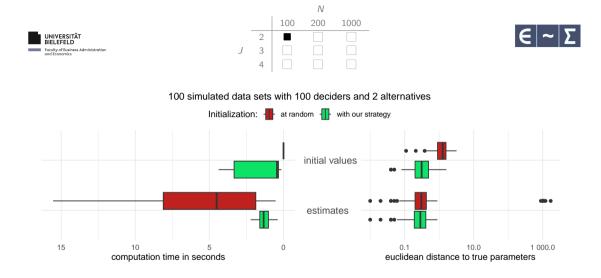


1 The multinomial probit model: purpose and estimation

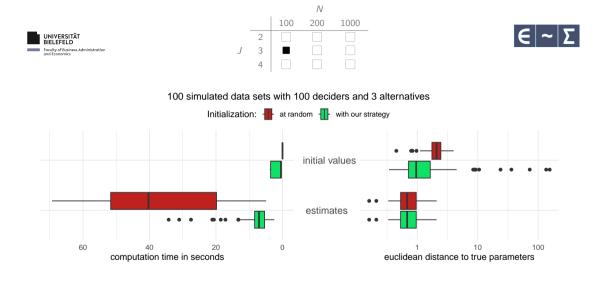
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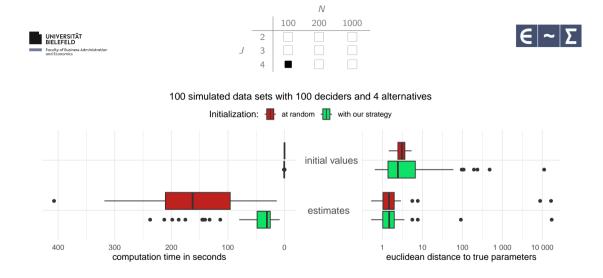


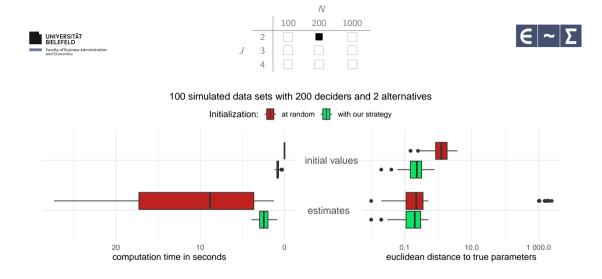


Simulation details: one alternative-varying regressor connected to alternative-varying coefficients; fully flexible Σ ; true parameters and random initial values were drawn from a standard normal distribution. Lennart Oelschläger Dietmar Bauer | Facilitating probit likelihood optimization

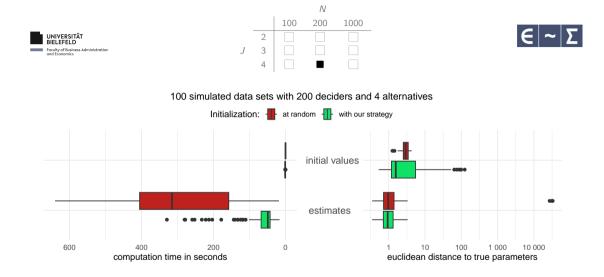


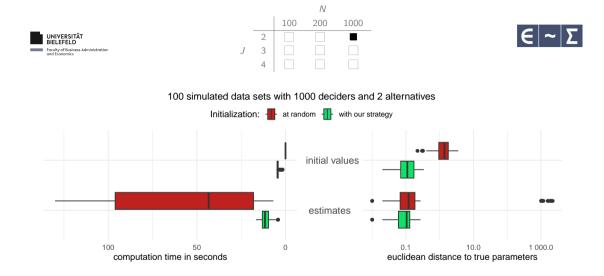
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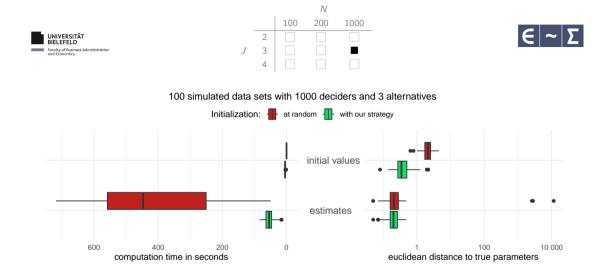


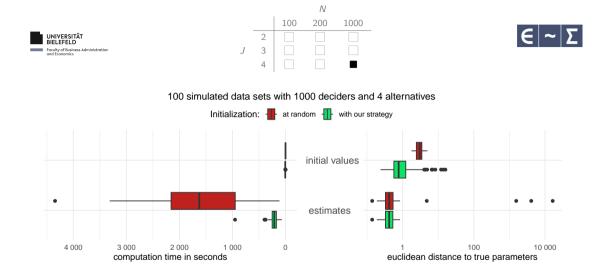
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- Our initialization idea improves optimization time and convergence rate.







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Thanks for your attention! Do you have any questions or comments?

lennart.oelschlaeger@uni-bielefeld.de
loelschlaeger.de/talks